

Large Dimensional Data in Econometrics: The Curses and Blessings

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 - Test statistic is not consistent, i.e., the power will not go to ∞ under the alternative

Example

Climate studies: T might be the number of time points and n the number of observation stations. $\frac{n}{T}$ is moderate.

Example

Financial data: large covariance and correlation matrices, with $n \approx 400$ financial indicators are publicly posted daily and used for value-at-risk calculations.

Example

Information Retrieval/search engines: A common search engine strategy forms huge term by document incidence matrices (n and T at least in the thousands).

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- Still OK, when $\frac{n}{T} \rightarrow 0$.
- However, when n is large and comparable with T , $\frac{n}{T} \rightarrow c$, it is known from RMT that the sample covariance matrix is no longer a good approximation to the population covariance matrix.

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- Spurious Regression: Phillips (1986). OLS of a spurious regression is inconsistent

$$y_t = \beta x_t + u_t$$

where $y_t \sim I(1)$, $x_t \sim I(1)$, and $u_t \sim I(1)$. y_t and x_t are independent. $\beta = 0$.

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- Spuriousness in time series can be removed in panel model with large n and large T .

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- $$\hat{\beta} = \frac{\sum_{i=1}^n \sum_{t=1}^T (x_t - \bar{x}) y_t}{\sum_{i=1}^n \sum_{t=1}^T (x_t - \bar{x})^2}.$$

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- $\hat{\beta} = \frac{\sum_{i=1}^n \sum_{t=1}^T (x_t - \bar{x}) y_t}{\sum_{i=1}^n \sum_{t=1}^T (x_t - \bar{x})^2}$.
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- $\int_0^1 E [W_x^2(r)] dr = \int_0^1 r dr = \frac{1}{2}, \int_0^1 E [\underline{W}_x^2(r)] dr = \frac{1}{6}$

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- $\int_0^1 E [W_x^2(r)] dr = \int_0^1 r dr = \frac{1}{2}, \int_0^1 E [\underline{W}_x^2(r)] dr = \frac{1}{6}$
- $\hat{\beta} \xrightarrow{p} 0$ as $T \rightarrow \infty$ and follows $n \rightarrow \infty$ sequentially.

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- $\hat{\beta}$ is consistent and the panel spurious model is identified.

- $$\sqrt{n}\hat{\beta} = \frac{\frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{1}{T^2} \sum_{t=1}^T (x_t - \bar{x}) y_t}{\frac{1}{n} \sum_{i=1}^n \frac{1}{T^2} \sum_{t=1}^T (x_t - \bar{x})^2}.$$

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- $\frac{1}{n} \sum_{i=1}^n \frac{1}{T^2} \sum_{t=1}^T (x_t - \bar{x})^2 \xrightarrow{p} \frac{\sigma_x^2}{6}$ as $(n, T) \rightarrow \infty$ sequentially/jointly.

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- $E \left[\int_0^1 \underline{B}_x(r) \underline{B}_y(r) dr \right] = 0$ and $Var \left[\int_0^1 \underline{B}_x(r) \underline{B}_y(r) dr \right] = \frac{\sigma_x^2 \sigma_y^2}{90}$

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- $\frac{1}{\sqrt{n}} \sum_{i=1}^n \int_0^1 \underline{B}_x(r) \underline{B}_y(r) dr \xrightarrow{d} N \left(0, \frac{2}{5} \frac{\sigma_y^2}{\sigma_x^2} \right)$ as $n \rightarrow \infty$

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- $\sqrt{n}\hat{\beta} = O_p(1)$.

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- Expand arsenal of econometric tools for thinking about these large panel models.

Factor-Loading Model

$$R_n = \Lambda F' + \varepsilon$$

where R_n is a $n \times T$ matrix of asset returns, F is an $T \times k$ matrix of factors, Λ is a $n \times k$ matrix of loadings and ε is a $n \times T$ matrix of idiosyncratic errors.

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- The sample eigenvalues are more spread out than the population eigenvalues.



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- What happens if $\frac{n}{T} \rightarrow c$?
- Test for constant conditional correlation (CCC) model.

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- Consistent Estimation with Weak Instruments in Panel Data

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 $T(\hat{\tau} - \tau_0) = O_p\left(\frac{1}{\|\delta_T\|^2}\right)$
- In time series need to assume: shrinking Break: $\delta_T \rightarrow 0$ and $T^{1/2-\alpha}\delta_T \rightarrow \infty$ for some $\alpha \in (0, \frac{1}{2})$

Change Point estimation in Panel Set-up

- $\hat{k} = k_0 + O_p\left(\frac{1}{n\|\delta_T\|^2}\right)$ \hat{k} is n -consistent $\delta_T = \beta_2 - \beta_1$

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$$\begin{pmatrix} \cdot \\ \vdots \\ \circ \\ \vdots \\ \cdot \end{pmatrix} \begin{matrix} 1 \\ \\ k_0 \\ \\ T \end{matrix}$$

a) Time series data

$$\begin{pmatrix} & & \cdot & & \\ & & \cdot & & \\ & & \cdot & & \\ \cdot & \dots & \circ & \dots & \cdot \\ & & \cdot & & \\ & & \cdot & & \\ & & \cdot & & \\ 1 & & i & & n \end{pmatrix} \begin{matrix} 1 \\ \\ k_0 \\ \\ T \end{matrix}$$

b) Panel data

Test for Cross-Sectional Dependence in Panel Models

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- Johnstone, I. M. (2001), "On the Distribution of the Largest Eigenvalue in Principal Component Analysis," *Annals of Statistics*, 29, 295-327.

Largest Eigenvalue of Random Covariance Matrices

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- We are looking for k vectors (e_1, e_2, \dots, e_k) in R^n such that $\sum_{m=1}^k \text{var}(\langle x_i, e_m \rangle)$ is maximal.

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- Principal component analysis (PCA). We are interested in recovering as much of the total variance in the data as possible while reducing the dimensionality of the problem from n to k .
- We are looking for k vectors (e_1, e_2, \dots, e_k) in R^n such that $\sum_{m=1}^k \text{var}(\langle x_i, e_m \rangle)$ is maximal.
- One should choose for e the eigenvectors associated with the first k eigenvalues $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_k$ of Σ_n .

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- Assume $\Sigma_n = I_n$. So all the population eigenvalues are equal to 1.

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- l_1 is an inconsistent estimator in the large n and T setting.